



ESTIMATION ERROR AND WHAT TO DO ABOUT IT

NORTHFIELD'S 29TH ANNUAL RESEARCH CONFERENCE

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Traditional approaches to estimation error

1. Equal Weighting

This approach avoids estimation altogether and builds an equally weighted portfolio.

2. Bayesian shrinkage

This approach compresses estimates toward a prior belief such as the cross-sectional average.

3. Resampling

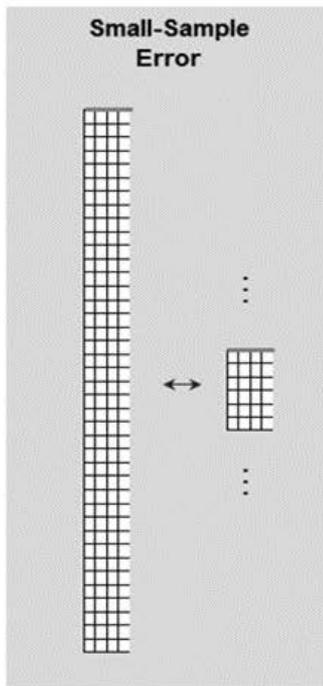
This approach repeatedly draws random samples from the data, generates efficient portfolios for each sample, and averages the weights for the portfolios at a chosen risk level.

4. Robust optimization

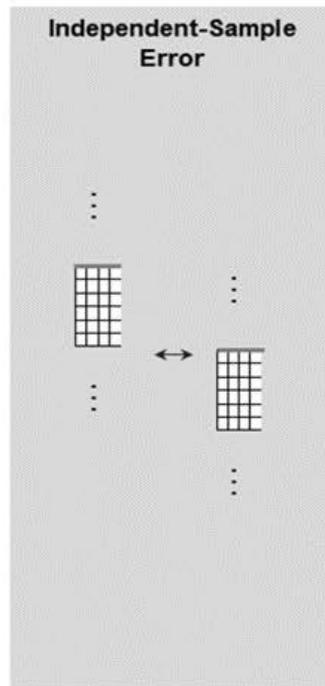
This approach considers a wide range of expected returns and risk and selects the portfolio that suffers the least in the most adverse scenario.



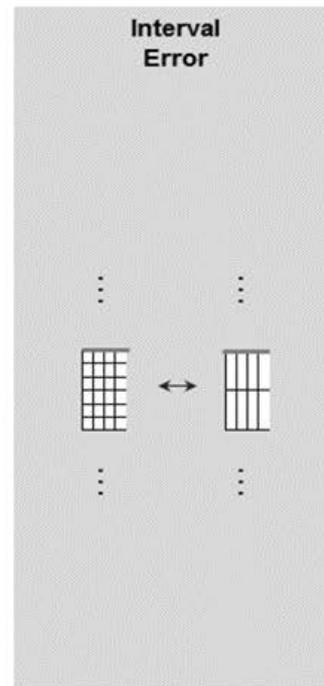
Types of estimation error



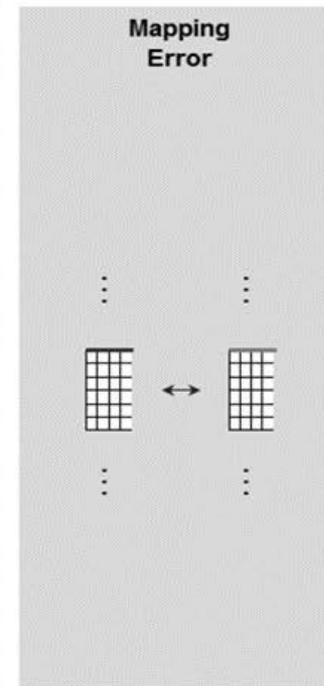
Vary: **Small samples**
Hold constant:
Forecasting sample
Factor mapping
Measurement interval



Vary: **Forecasting sample**
Hold constant:
Sample size
Factor mapping
Measurement interval



Vary: **Measurement interval**
Hold constant:
Sample size
Forecasting sample
Factor mapping



Vary: **Factor mapping**
Hold constant:
Sample size
Forecasting sample
Measurement interval



Small-sample error

The realization of parameters from a small sample will likely differ from the parameter values of a large sample from which it is selected.

We call this small-sample error.

$$SSE(A, B) = \sqrt{\frac{1}{n} \sum_{j=1}^n \left(\frac{\rho_{AB,m,j} \sqrt{\sigma_{A,m,j} \sigma_{B,m,j}} - \rho_{AB,m} \sqrt{\sigma_{A,m} \sigma_{B,m}}}{\sqrt{\sigma_{A,m} \sigma_{B,m}}} \right)^2}$$

When A and B are the same asset, this formula will measure the error in the standard deviation of that asset.



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m, j indicates monthly estimates
from a 36-month testing subsample

m alone indicates monthly estimates
from the full sample

When A and B are the same asset, this formula will measure the error in the standard deviation of that asset.



Independent-sample error

The realization of parameters from a future sample will likely differ from the parameter values of an independent historical sample.

We call this independent-sample error.

$$ISE(A, B) = \sqrt{\frac{1}{n} \sum_{j=1}^n \left(\frac{\rho_{AB,m,j} \sqrt{\sigma_{A,m,j} \sigma_{B,m,j}} - \rho_{AB,\hat{m},j} \sqrt{\sigma_{A,\hat{m},j} \sigma_{B,\hat{m},j}}}{\sqrt{\sigma_{A,m} \sigma_{B,m}}} \right)^2 - SSE(A, B)^2}$$



Independent-sample error

The realization of parameters from a future sample will likely differ from the parameter values of an independent historical sample.

We call this independent-sample error.

$$ISE(A, B) = \sqrt{\frac{1}{n} \sum_{j=1}^n \left(\frac{\rho_{AB,m,j} \sqrt{\sigma_{A,m,j} \sigma_{B,m,j}} - \rho_{AB,\hat{m},j} \sqrt{\sigma_{A,\hat{m},j} \sigma_{B,\hat{m},j}}}{\sqrt{\sigma_{A,m} \sigma_{B,m}}} \right)^2 - SSE(A, B)^2}$$

m, j indicates monthly estimates
from a 36-month testing subsample

\hat{m}, j indicates monthly estimates
from a 36-month independent
subsample immediately preceding
the testing subsample



Mapping error

Assets define the opportunity set for investing. A desired factor exposure must be mapped onto a portfolio of assets to be investable.

The mapping that best tracks a factor in the future will likely differ from the mapping of an independent historical sample.

We call this mapping error.

Mapping error applies only to factors.

The calculations for small sample error and independent sample error for factors do not reflect mapping error.



Mapping error

Mapping error can be isolated by comparing the covariance of the best-fit factor-mimicking portfolio to the covariance of a factor-mimicking portfolio estimated from an independent sample, holding the evaluation period constant.

$$ME(A, B) = \sqrt{\frac{1}{n} \sum_{j=1}^n \left(\frac{\rho_{\hat{A}\hat{B},m,j} \sqrt{\sigma_{\hat{A},m,j} \sigma_{\hat{B},m,j}} - \rho_{AB,m,j} \sqrt{\sigma_{A,m,j} \sigma_{B,m,j}}}{\sqrt{\sigma_{A,m} \sigma_{B,m}}} \right)^2}$$



Mapping error

Mapping error can be isolated by comparing the covariance of the best-fit factor-mimicking portfolio to the covariance of a factor-mimicking portfolio estimated from an independent sample, holding the evaluation period constant.

$$ME(A, B) = \sqrt{\frac{1}{n} \sum_{j=1}^n \left(\frac{\rho_{\hat{A}\hat{B},m,j} \sqrt{\sigma_{\hat{A},m,j} \sigma_{\hat{B},m,j}} - \rho_{AB,m,j} \sqrt{\sigma_{A,m,j} \sigma_{B,m,j}}}{\sqrt{\sigma_{A,m} \sigma_{B,m}}} \right)^2}$$



\hat{A}, \hat{B} indicate factor mappings derived from the independent subsample



A, B indicate factor mappings derived from the testing subsample



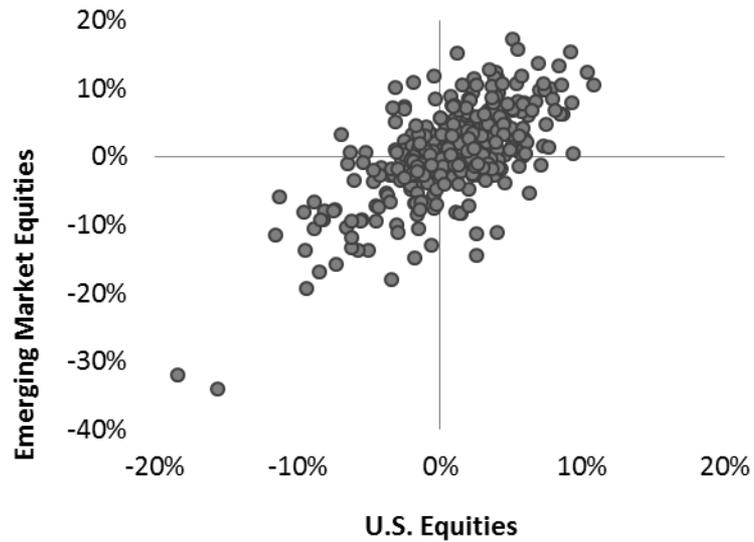
Interval error

Parameters estimated from monthly or higher-frequency returns often differ from those estimated from lower-frequency returns, within the same sample.

We call this interval error.



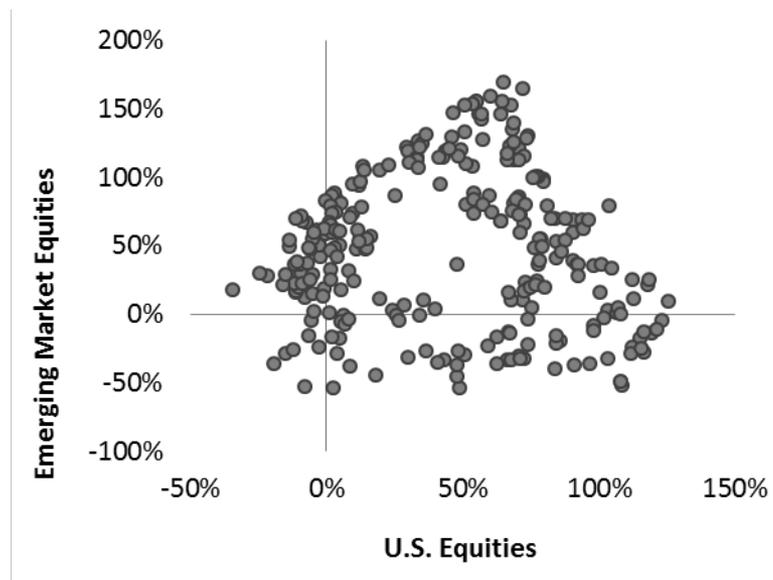
U.S. and Emerging Markets Stocks:
Monthly Returns - January 1988 through December 2015



Correlation = 0.68



U.S. and Emerging Markets Stocks:
Five-Year Returns - January 1988 through December 2015



Correlation = -0.04



Interval error: Long-horizon and short-horizon volatility

The volatility of the cumulative continuous returns of x over q periods is given by:

$$\sigma(x_t + \dots + x_{t+q-1}) = \sigma_x \sqrt{q + 2 \sum_{k=1}^{q-1} (q-k) \rho_{x_t, x_{t+k}}}$$



Interval error: Long-horizon and short-horizon volatility

The volatility of the cumulative continuous returns of x over q periods is given by:

$$\sigma(x_t + \dots + x_{t+q-1}) = \sigma_x \sqrt{q + 2 \sum_{k=1}^{q-1} (q-k) \rho_{x_t, x_{t+k}}}$$



This term reflects annualization in the absence of lagged effects



Interval error: Long-horizon and short-horizon volatility

The volatility of the cumulative continuous returns of x over q periods is given by:

$$\sigma(x_t + \dots + x_{t+q-1}) = \sigma_x \sqrt{q + 2 \sum_{k=1}^{q-1} (q-k) \rho_{x_t, x_{t+k}}}$$



This term captures the impact of auto-correlation



Interval error: Long-horizon and short-horizon correlation

The correlation between the cumulative returns of x and the cumulative returns of y over q periods is given by:

$$\rho(x_t + \dots + x_{t+q-1}, y_t + \dots + y_{t+q-1}) = \frac{q\rho_{x_t, y_t} + \sum_{k=1}^{q-1} (q-k)(\rho_{x_{t+k}, y_t} + \rho_{x_t, y_{t+k}})}{\sqrt{q + 2\sum_{k=1}^{q-1} (q-k)\rho_{x_t, x_{t+k}}} \sqrt{q + 2\sum_{k=1}^{q-1} (q-k)\rho_{y_t, y_{t+k}}}}$$



Interval error: Long-horizon and short-horizon correlation

The correlation between the cumulative returns of x and the cumulative returns of y over q periods is given by:

$$\rho(x_t + \dots + x_{t+q-1}, y_t + \dots + y_{t+q-1}) = \frac{q\rho_{x_t, y_t} + \sum_{k=1}^{q-1} (q-k)(\rho_{x_{t+k}, y_t} + \rho_{x_t, y_{t+k}})}{\sqrt{q + 2\sum_{k=1}^{q-1} (q-k)\rho_{x_t, x_{t+k}}} \sqrt{q + 2\sum_{k=1}^{q-1} (q-k)\rho_{y_t, y_{t+k}}}}$$

This term captures the lagged cross-correlation between x and y



Interval error: Long-horizon and short-horizon correlation

The correlation between the cumulative returns of x and the cumulative returns of y over q periods is given by:

$$\rho(x_t + \dots + x_{t+q-1}, y_t + \dots + y_{t+q-1}) = \frac{q\rho_{x_t, y_t} + \sum_{k=1}^{q-1} (q-k)(\rho_{x_{t+k}, y_t} + \rho_{x_t, y_{t+k}})}{\sqrt{q + 2\sum_{k=1}^{q-1} (q-k)\rho_{x_t, x_{t+k}}} \sqrt{q + 2\sum_{k=1}^{q-1} (q-k)\rho_{y_t, y_{t+k}}}}$$

↑
This term captures the auto-correlation of x

↑
This term captures the auto-correlation of y



Interval error

Interval error can be isolated by comparing a covariance matrix estimated from low-frequency returns to a covariance matrix estimated from high-frequency returns.

$$IE(A, B) = \sqrt{\frac{1}{n} \sum_{j=1}^n \left(\frac{\rho_{AB,ann,j} \sqrt{\sigma_{A,ann,j} \sigma_{B,ann,j}} / 12 - \rho_{AB,m,j} \sqrt{\sigma_{A,m,j} \sigma_{B,m,j}}}{\sqrt{\sigma_{A,m} \sigma_{B,m}}} \right)^2}$$



Interval error

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$$IE(A, B) = \sqrt{\frac{1}{n} \sum_{j=1}^n \left(\frac{\rho_{AB,ann,j} \sqrt{\sigma_{A,ann,j} \sigma_{B,ann,j}} / 12 - \rho_{AB,m,j} \sqrt{\sigma_{A,m,j} \sigma_{B,m,j}}}{\sqrt{\sigma_{A,m} \sigma_{B,m}}} \right)^2}$$



tri, j indicates implied 3-year estimates from a 36-month testing subsample



m, j indicates monthly estimates from the same 36-month testing sample



Composite instability score

The four sources of estimation error are independent from one another, which means we can sum the variances of each error and then take the square root of this sum to compute a composite instability score.

$$\text{Composite Instability Score (CIS)} = \sqrt{SSE^2 + ISE^2 + ME^2 + IE^2}$$



Stability-adjusted return distribution

Begin with a long sample of asset returns



Estimate small sample covariance matrices

$$\Sigma_{s1}$$

$$\Sigma_{s2}$$

$$\Sigma_{s3}$$

...

$$\Sigma_{sn}$$

Compute error matrix for each small sample versus its complementary sample

$$\Sigma_{e1}$$

$$\Sigma_{e2}$$

$$\Sigma_{e3}$$

...

$$\Sigma_{en}$$

$$= \Sigma_{s1} - \Sigma_{c1}$$

$$= \Sigma_{s2} - \Sigma_{c2}$$

$$= \Sigma_{s3} - \Sigma_{c3}$$

$$= \Sigma_{sn} - \Sigma_{cn}$$

Add each error matrix to the baseline covariance matrix

$$\Sigma_1$$

$$\Sigma_2$$

$$\Sigma_3$$

...

$$\Sigma_n$$

$$= \Sigma_{e1} + \Sigma_{base}$$

$$= \Sigma_{e2} + \Sigma_{base}$$

$$= \Sigma_{e3} + \Sigma_{base}$$

$$= \Sigma_{en} + \Sigma_{base}$$

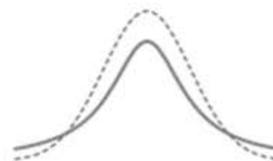
Draw sample returns from normal distributions



...



Combine to form a composite non-normal distribution





Implications for portfolio construction

- Mean-variance analysis assumes either that investors have preferences that can be well approximated by mean and variance or that returns are elliptically distributed.
- If sub-samples within a large sample have different correlations or if kurtosis is not uniform across asset classes, the return distribution will not be elliptical.
- Some investors face thresholds and therefore have preferences that are better represented by a kinked utility function which displays sharp aversion to losses below a given threshold. Mean and variance are not sufficient to describe kinked utility functions.
- If returns are not elliptical and investors have preferences that cannot be approximated by mean and variance, it may be preferable to employ full-scale optimization to identify the optimal portfolio.



Implications for portfolio construction

Full-scale optimization: Illustrative example

	Stock Return	Bond Return	Stock Weight	Bond Weight	Utility Calculation	Utility
1993	10.1%	16.2%	100%	0%	$\text{Ln} [(1 + 10.1\%) \times 100\% + (1 + 16.2\%) \times 0\%] =$	0.0962
1994	1.3%	-7.1%	100%	0%	$\text{Ln} [(1 + 1.3\%) \times 100\% + (1 - 7.1\%) \times 0\%] =$	0.0129
1995	37.5%	30.0%	100%	0%	$\text{Ln} [(1 + 39.5\%) \times 100\% + (1 + 30.0\%) \times 0\%] =$	0.3185
1996	22.9%	0.1%	100%	0%	$\text{Ln} [(1 + 22.9\%) \times 100\% + (1 + 0.1\%) \times 0\%] =$	0.2062
1997	33.3%	14.5%	100%	0%	$\text{Ln} [(1 + 33.3\%) \times 100\% + (1 + 14.5\%) \times 0\%] =$	0.2874
1998	28.6%	11.8%	100%	0%	$\text{Ln} [(1 + 28.6\%) \times 100\% + (1 + 11.8\%) \times 0\%] =$	0.2515
1999	20.9%	-7.6%	100%	0%	$\text{Ln} [(1 + 20.9\%) \times 100\% + (1 - 7.6\%) \times 0\%] =$	0.1898
2000	-9.1%	16.1%	100%	0%	$\text{Ln} [(1 - 9.1\%) \times 100\% + (1 + 16.2\%) \times 0\%] =$	-0.0954
2001	-11.9%	7.3%	100%	0%	$\text{Ln} [(1 - 11.9\%) \times 100\% + (1 + 7.3\%) \times 0\%] =$	-0.1267
2002	-22.1%	14.8%	100%	0%	$\text{Ln} [(1 - 22.1\%) \times 100\% + (1 + 14.8\%) \times 0\%] =$	-0.2497

Average utility = **0.0891**



Implications for portfolio construction

Full-scale optimization: Illustrative example

	Stock Return	Bond Return	Stock Weight	Bond Weight	Utility Calculation	Utility
1993	10.1%	16.2%	50%	50%	$\text{Ln} [(1 + 10.1\%) \times 50\% + (1 + 16.2\%) \times 50\%] =$	0.1235
1994	1.3%	-7.1%	50%	50%	$\text{Ln} [(1 + 1.3\%) \times 50\% + (1 - 7.1\%) \times 50\%] =$	-0.0294
1995	37.5%	30.0%	50%	50%	$\text{Ln} [(1 + 39.5\%) \times 50\% + (1 + 30.0\%) \times 50\%] =$	0.2908
1996	22.9%	0.1%	50%	50%	$\text{Ln} [(1 + 22.9\%) \times 50\% + (1 + 0.1\%) \times 50\%] =$	0.1089
1997	33.3%	14.5%	50%	50%	$\text{Ln} [(1 + 33.3\%) \times 50\% + (1 + 14.5\%) \times 50\%] =$	0.2143
1998	28.6%	11.8%	50%	50%	$\text{Ln} [(1 + 28.6\%) \times 50\% + (1 + 11.8\%) \times 50\%] =$	0.1840
1999	20.9%	-7.6%	50%	50%	$\text{Ln} [(1 + 20.9\%) \times 50\% + (1 - 7.6\%) \times 50\%] =$	0.0644
2000	-9.1%	16.1%	50%	50%	$\text{Ln} [(1 - 9.1\%) \times 50\% + (1 + 16.2\%) \times 50\%] =$	0.0344
2001	-11.9%	7.3%	50%	50%	$\text{Ln} [(1 - 11.9\%) \times 50\% + (1 + 7.3\%) \times 50\%] =$	-0.0233
2002	-22.1%	14.8%	50%	50%	$\text{Ln} [(1 - 22.1\%) \times 50\% + (1 + 14.8\%) \times 50\%] =$	-0.0372

Average utility = **0.0930**



Implications for portfolio construction

Full-scale optimization: Illustrative example

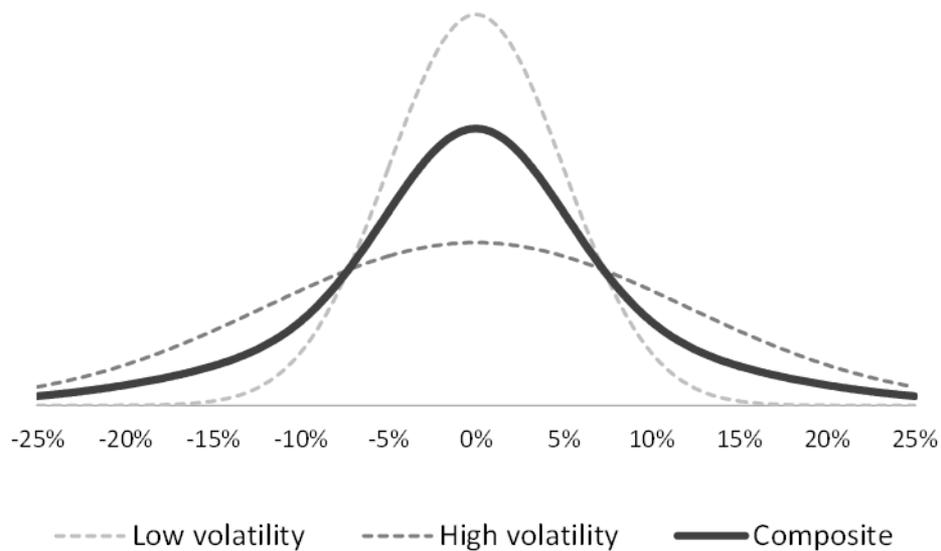
	Stock Return	Bond Return	Stock Weight	Bond Weight	Utility Calculation	Utility
1993	10.1%	16.2%	55%	45%	$\text{Ln} [(1 + 10.1\%) \times 55\% + (1 + 16.2\%) \times 45\%] =$	0.1208
1994	1.3%	-7.1%	55%	45%	$\text{Ln} [(1 + 1.3\%) \times 55\% + (1 - 7.1\%) \times 45\%] =$	-0.0251
1995	37.5%	30.0%	55%	45%	$\text{Ln} [(1 + 39.5\%) \times 55\% + (1 + 30.0\%) \times 45\%] =$	0.2936
1996	22.9%	0.1%	55%	45%	$\text{Ln} [(1 + 22.9\%) \times 55\% + (1 + 0.1\%) \times 45\%] =$	0.1190
1997	33.3%	14.5%	55%	45%	$\text{Ln} [(1 + 33.3\%) \times 55\% + (1 + 14.5\%) \times 45\%] =$	0.2219
1998	28.6%	11.8%	55%	45%	$\text{Ln} [(1 + 28.6\%) \times 55\% + (1 + 11.8\%) \times 45\%] =$	0.1910
1999	20.9%	-7.6%	55%	45%	$\text{Ln} [(1 + 20.9\%) \times 55\% + (1 - 7.6\%) \times 45\%] =$	0.0777
2000	-9.1%	16.1%	55%	45%	$\text{Ln} [(1 - 9.1\%) \times 55\% + (1 + 16.2\%) \times 45\%] =$	0.0222
2001	-11.9%	7.3%	55%	45%	$\text{Ln} [(1 - 11.9\%) \times 55\% + (1 + 7.3\%) \times 45\%] =$	-0.0331
2002	-22.1%	14.8%	55%	45%	$\text{Ln} [(1 - 22.1\%) \times 55\% + (1 + 14.8\%) \times 45\%] =$	-0.0565

Average utility =

0.0931

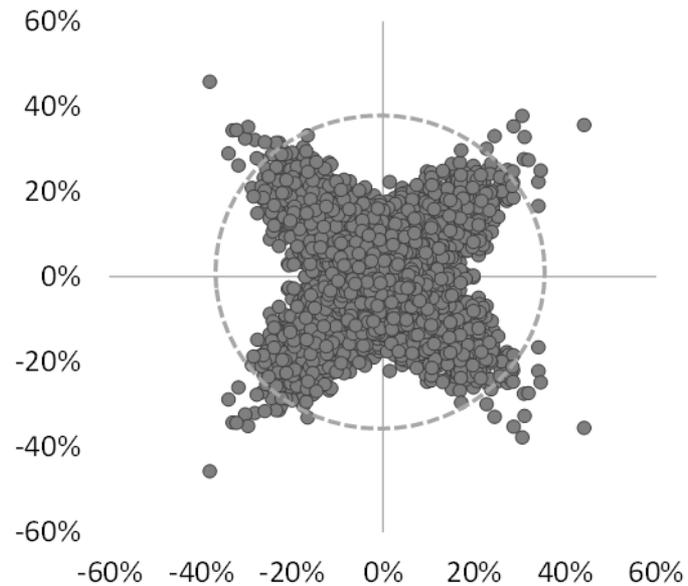


Mixture of Two Normal Distributions





Multivariate Mixture of Asset Classes with Unstable Correlation





Asset allocation examples - methodology

- We select the first five-year subsample from the full sample and set it aside.
- We then build a stability-adjusted return sample using the remaining data in the original sample.
- Next we use this complementary sample to build six portfolios: a portfolio that ignores errors, one that applies Bayesian shrinkage, and one formed from the stability-adjusted return sample, all using full-scale optimization, and then again using mean-variance analysis.
- We repeat steps 1 through 3 for all 36 testing samples, which are overlapping periods ending in December.
- Using a variety of metrics, we evaluate each portfolio in the subsample that was held out of the complementary sample used to form it.
- We shrink the standard deviations by blending them equally with their cross-sectional mean, and we do the same for the correlations.



Asset allocation using full-scale optimization

Average Optimal Weights	Ignoring Errors	Bayesian Shrinkage	Stability Adjusted
U.S. Equities	25.0%	20.1%	33.3%
Foreign Developed Market Equities	19.6%	18.9%	16.7%
Emerging Market Equities	13.6%	17.2%	11.8%
Treasury Bonds	20.1%	13.9%	27.9%
U.S. Corporate Bonds	15.6%	15.3%	7.9%
Commodities	3.8%	5.7%	1.4%
Cash Equivalents	2.4%	8.9%	1.0%
10 Percentile Worst Outcome Across Testing Samples			
12-Month Volatility	16.9%	16.9%	15.8%
12-Month Value at Risk (10% significance)	-24.1%	-23.9%	-22.9%
12-Month Value at Risk (5% significance)	-28.5%	-28.3%	-25.2%
Worst 12-Month Return	-32.7%	-32.5%	-29.3%

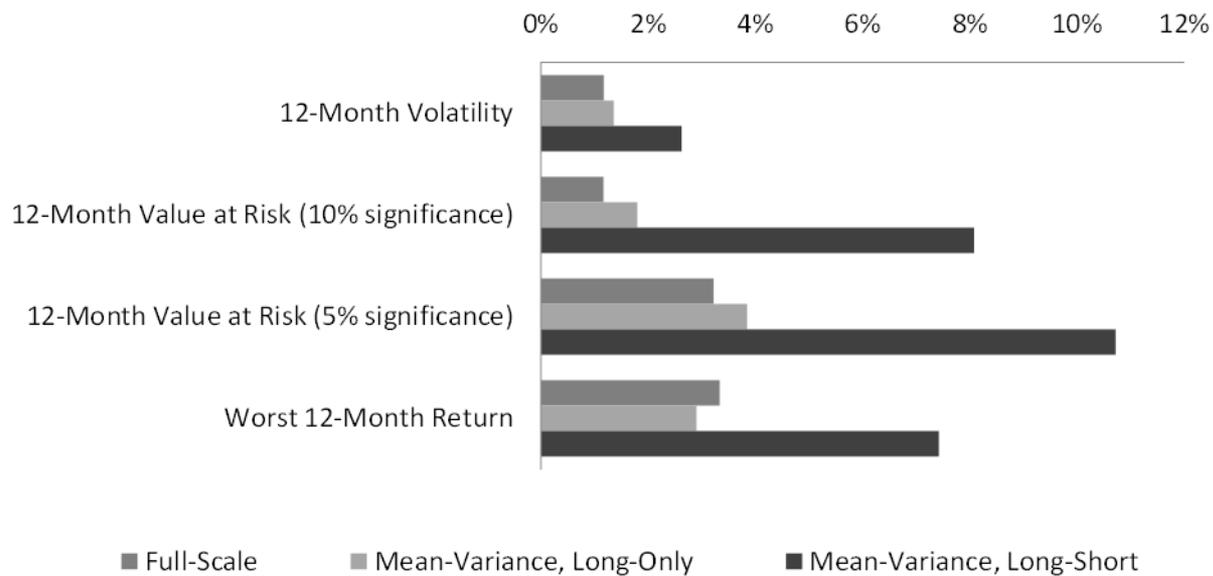


Asset allocation using mean-variance analysis

Average Optimal Weights	Long Only			Long-Short		
	Ignoring Errors	Bayesian Shrinkage	Stability Adjusted	Ignoring Errors	Bayesian Shrinkage	Stability Adjusted
U.S. Equities	24.7%	21.2%	32.5%	24.6%	21.2%	33.4%
Foreign Developed Market Equities	19.4%	18.8%	18.0%	19.3%	18.8%	20.2%
Emerging Market Equities	14.1%	16.7%	10.9%	13.8%	16.7%	13.0%
Treasury Bonds	22.4%	13.7%	24.6%	30.7%	13.7%	56.7%
U.S. Corporate Bonds	13.9%	14.8%	12.3%	12.2%	14.8%	-13.1%
Commodities	3.6%	5.9%	1.3%	4.0%	5.9%	-5.4%
Cash Equivalents	1.9%	9.0%	0.3%	-4.7%	9.0%	-4.8%
10 Percentile Worst Outcome Across Testing Samples						
12-Month Volatility	17.1%	17.1%	15.8%	18.7%	17.1%	16.1%
12-Month Value at Risk (10% significance)	-25.3%	-24.9%	-23.5%	-29.9%	-24.9%	-21.8%
12-Month Value at Risk (5% significance)	-29.7%	-28.9%	-25.9%	-34.6%	-28.9%	-23.9%
Worst 12-Month Return	-33.0%	-32.6%	-30.1%	-35.7%	-32.6%	-28.2%



Stability-adjusted benefit compared to error-blind optimization





Summary



- When investors estimate covariances from historical returns they face four types of estimation error: small-sample error, independent-sample error, mapping error, and interval error.
- Small-sample error arises because the investor's investment horizon is typically shorter than the historical sample from which covariances are estimated.
- Independent-sample error arises because the investor's investment horizon is independent of history.
- Interval error arises because investors estimate covariances from higher frequency returns than the return frequency they care about. If returns have non-zero auto-correlations, standard deviation does not scale with the square root of time. If returns have non-zero auto-correlations or non-zero lagged cross correlations, correlation is not invariant to the return interval used to measure it.
- Common approaches for controlling estimation error, such as Bayesian shrinkage and resampling, make portfolios less sensitive to estimation error.
- A new approach, called stability-adjusted optimization, assumes that some covariances are reliably more stable than other covariances. It delivers portfolios that rely more on relatively stable covariances and less on relatively unstable covariances.